# Consistent Risk Measures and a Non-linear Extension of Backwards Martingale Convergence

Hans Föllmer<sup>\*</sup> Irina Penner<sup>†</sup>

August 17, 2014

#### Abstract

We study the behavior of conditional risk measures along decreasing  $\sigma$ -fields. Under a condition of consistency, we prove a non-linear extension of backwards martingale convergence. In particular we show the existence of a limiting conditional risk measure with respect to the tail field, we describe its dual representation in terms of a limiting penalty function, and we show that consistency extends to the tail field. Moreover, we clarify the structure of global risk measures which are consistent with the given sequence of conditional risk measures.

### 1 Introduction

Consider a filtration  $(\mathcal{F}_n)_{n \in \mathbb{Z}}$ , indexed by the integers, on some measurable space  $(\Omega, \mathcal{F})$ . In the forward direction we define the asymptotic  $\sigma$ -field  $\mathcal{F}_{\infty} := \sigma (\bigcup_n \mathcal{F}_n)$ , in the backward direction the tail field  $\mathcal{F}_{-\infty} := \bigcap_n \mathcal{F}_n$ .

For a given probability measure P and for any bounded measurable function X on  $(\Omega, \mathcal{F})$ , let us denote by

$$\eta_n(X) := E_P\left[-X|\mathcal{F}_n\right], \qquad n \in \mathbb{Z}$$
(1)

the conditional expectation of -X with respect to  $\mathcal{F}_n$  under the measure P. Since we are using the minus sign, the functional  $\eta_n$  can be regarded as the special linear case of a conditional convex risk measure, as explained below.

Due to the projectivity of conditional expectations, the sequence  $(\eta_n)_{n\in\mathbb{Z}}$  is *consistent* in the sense that

$$\eta_n(-\eta_{n+1}(X)) = \eta_n(X), \qquad n \in \mathbb{Z}.$$
(2)

<sup>\*</sup>Humboldt-Universität zu Berlin, Institut für Mathematik, Unter den Linden 6, 10099 Berlin, Germany. Email: foellmer@math.hu-berlin.de.

<sup>&</sup>lt;sup>†</sup>Humboldt-Universität zu Berlin, Institut für Mathematik, Unter den Linden 6, 10099 Berlin, Germany. Email: penner@math.hu-berlin.de. Supported by the DFG Research Center MATHEON "Mathematics for key technologies".

Martingale convergence forwards and backwards yields the existence of the limits

$$\eta_{\infty}(X) := \lim_{n \uparrow \infty} \eta_n(X), \quad \eta_{-\infty}(X) := \lim_{n \downarrow -\infty} \eta_n(X)$$

*P*-a.s. and in  $L^{1}(P)$ , and these limits are identified as conditional expectations

$$\eta_{\infty}(X) = E_P\left[-X|\mathcal{F}_{\infty}\right], \quad \eta_{-\infty}(X) = E_P\left[-X|\mathcal{F}_{-\infty}\right]$$

with respect to the limiting  $\sigma$ -fields  $\mathcal{F}_{\infty}$  and  $\mathcal{F}_{-\infty}$ . Again by projectivity, we see that the consistency relation (2) extends to infinity in both directions, that is,

$$\eta_n(-\eta_\infty) = \eta_n \quad \text{and} \quad \eta_{-\infty}(-\eta_n) = \eta_{-\infty}$$
(3)

for any  $n \in \mathbb{Z}$ . Let us summarize these classical facts by saying that the sequence  $(\eta_n)_{n \in \mathbb{Z}}$  is asymptotically precise in both directions.

In this paper, we study the question whether asymptotic precision extends from the linear case of conditional expectation to the non-linear case of conditional risk measures. For each  $n \in \mathbb{Z}$ , let  $\rho_n$  denote a conditional convex risk measure on  $L^{\infty}(\Omega, \mathcal{F}, P)$  with respect to  $\mathcal{F}_n$ , and let

$$\mathcal{A}_n := \left\{ X \in L^{\infty}(\Omega, \mathcal{F}, P) \mid \rho_n(X) \le 0 \right\}$$

denote the corresponding acceptance set; see, e.g., [19, Chapter 11]. Under an additional continuity assumption, the conditional risk measure  $\rho_n$  admits the dual representation

$$\rho_n(X) = \underset{\substack{Q \ll P\\Q \approx P \text{ on } \mathcal{F}_n}}{\operatorname{ess \, sup}} \left( E_Q \left[ -X | \mathcal{F}_n \right] - \alpha_n(Q) \right) \tag{4}$$

with penalty function

$$\alpha_n(Q) = \operatorname{ess\,sup}_{X \in \mathcal{A}_n} E_Q\left[-X|\mathcal{F}_n\right].$$

In the special coherent case where  $\rho_n$  is also positively homogeneous, this reduces to the representation

$$\rho_n(X) = \operatorname{ess\,sup}_{Q \in \mathcal{Q}_n} E_Q\left[-X|\mathcal{F}_n\right] \tag{5}$$

with a suitable class  $Q_n$  of probability measures Q. Under the additional condition of comonotonicity, the coherent risk measure in (5) can also be regarded as a conditional Choquet integral

$$\rho_n(X) = \int (-X) dC_n,$$

where  $C_n(A) := \rho_n(-I_A)$  is a conditional Choquet capacity, in analogy to the discussion in [19, Section 4.7]. Clearly, we recover the conditional expectation

 $\eta_n(X)$  in (1) in the simple special case, where the set  $\mathcal{Q}_n$  reduces to the single probability measure P.

Let us now assume that the sequence  $(\rho_n)_{n \in \mathbb{Z}}$  of conditional risk measures is consistent in the sense of (2), i.e.,

$$\rho_n(-\rho_{n+1}(X)) = \rho_n(X), \qquad n \in \mathbb{Z}$$
(6)

for any  $X \in L^{\infty}(\Omega, \mathcal{F}, P)$ . Consistency can be characterized in terms of the acceptance sets  $(\mathcal{A}_n)_{n \in \mathbb{Z}}$ , in terms of the penalty functions  $(\alpha_n)_{n \in \mathbb{Z}}$ , and also by supermartingale criteria for the joint behavior of  $(\rho_n)$  and  $(\alpha_n)$ ; this is recalled in Section 2.2.

In the forward direction, the behavior of the consistent sequence  $(\rho_n)$  along the filtration  $(\mathcal{F}_n)_{n\geq 0}$  has been studied in [17]. The supermartingale criteria for consistency yield existence of the limit

$$\rho_{\infty}(X) := \lim \rho_n(X).$$

The question is whether  $\rho_{\infty}$  has good properties as a conditional risk measure with respect to  $\mathcal{F}_{\infty}$ . In the case  $\mathcal{F}_{\infty} = \mathcal{F}$ , asymptotic precision in the forward direction amounts to the condition  $\rho_{\infty}(X) = -X$ . However, neither asymptotic precision nor the weaker condition  $\rho_{\infty}(X) \ge -X$  of asymptotic safety may hold; see [17, Section 5] for criteria and for counterexamples.

In this paper, we focus on the backward direction, and so it is enough to consider the filtration  $(\mathcal{F}_n)_{n\leq 0}$ . Under a mild condition on the penalties for our reference measure P, we show in Section 3 that asymptotic precision is indeed satisfied along decreasing  $\sigma$ -fields. More precisely, an application of the supermartingale criteria for consistency yields the existence of the limit

$$\rho_{-\infty}(X) = \lim_{n \downarrow -\infty} \rho_n(X).$$

We then show that the functional  $\rho_{-\infty}$  defines a conditional convex risk measure with respect to the tail field  $\mathcal{F}_{-\infty}$ , that this risk measure is continuous from above, and that its dual representation (4) for  $n = -\infty$  is given by the limiting penalty function

$$\alpha_{-\infty}(Q) = \lim_{n \to \infty} \alpha_n(Q).$$

Moreover, we show that the consistency condition (6) extends to  $-\infty$ , that is,

$$\rho_{-\infty}(-\rho_n) = \rho_{-\infty}$$

for any  $n \leq 0$ , in analogy to (3). In particular, these properties of asymptotic precision in the backward direction hold for a consistent sequence of conditional coherent risk measures, and also for the special case of conditional Choquet integrals.

In the final Section 4 we study the structure of the set  $\mathcal{R}$  of all global (unconditional) risk measures  $\rho$  on  $L^{\infty}(\Omega, \mathcal{F}, P)$ , which are consistent with the

given sequence  $(\rho_n)_{n \leq 0}$ . Under additional continuity conditions, we show that such risk measures are of the form

$$\rho = \hat{\rho}(-\rho_{-\infty}),$$

where  $\hat{\rho}$  is a convex risk measure on the tail field; the precise formulation is given in Theorem 5 and Corollary 5.

Our discussion of the behavior of conditional convex risk measures along decreasing  $\sigma$ -fields is motivated by the problem of clarifying the structure of spatial risk measures consistent with a given local specification in a large network. Under a condition of local law-invariance, the local conditional risk measures must be entropic, and then the problem can be solved explicitly, as shown in [15]. Without this condition, the main problem consists in extending the local specification to the tail-field, and this can be done by using the general convergence results of the present paper. The application to spatial risk measures will be discussed in [16].

### 2 Preliminaries

Throughout this paper we fix a probability space  $(\Omega, \mathcal{F}, P)$ . We write  $L^{\infty} := L^{\infty}(\Omega, \mathcal{F}, P)$  and denote by  $\mathcal{M}_1(P)$  the set of all probability measures absolutely continuous with respect to P.

In this section we recall some basic facts about conditional convex risk measures and about consistency that will be used later on. For further details see, for example, [13, 4, 17, 5, 8, 1], and [19, Chapter 11].

### 2.1 Conditional convex risk measures

Let  $\mathcal{F}_0$  be a sub- $\sigma$ -field of  $\mathcal{F}$  and write  $L_0^{\infty} := L^{\infty}(\Omega, \mathcal{F}, P)$ .

#### **Definition 1.** A map

$$\rho_0: L^\infty \to L_0^\infty$$

is called a conditional convex risk measure with respect to  $\mathcal{F}_0$  if it satisfies the following properties for any  $X, Y \in L^{\infty}$ :

• Conditional cash invariance: For all  $X_0 \in L_0^{\infty}$ ,

$$\rho_0(X + X_0) = \rho_0(X) - X_0$$

- Monotonicity:  $X \leq Y \Rightarrow \rho_0(X) \geq \rho_0(Y)$
- Conditional convexity: For all  $\lambda \in L_0^\infty$  such that  $0 \le \lambda \le 1$ ,

$$\rho_0(\lambda X + (1-\lambda)Y) \le \lambda \rho_0(X) + (1-\lambda)\rho_0(Y)$$

• Normalization:  $\rho_0(0) = 0$ .

A conditional convex risk measure  $\rho_0$  is called a conditional coherent risk measure if it has in addition the following property:

• Conditional positive homogeneity: For all  $\lambda \in L_0^{\infty}$  such that  $\lambda \ge 0$ ,

$$\rho_0(\lambda X) = \lambda \rho_0(X).$$

**Remark 1.** A conditional convex risk measure  $\rho_0$  is uniquely determined by the associated acceptance set

$$\mathcal{A}_0 := \left\{ X \in L^\infty \mid \rho_0(X) \le 0 \right\},\,$$

since

$$\rho_0(X) = \operatorname{ess\,inf}\left\{ Y \in L_0^\infty \mid X + Y \in \mathcal{A}_0 \right\}.$$
(7)

Thus  $\rho_n(X)$  has the financial interpretation of a capital requirement, namely the minimal amount which should be added to the position X to make it acceptable.

Note that  $\mathcal{A}_0$  is conditionally convex and solid, and that  $\rho_0(0) = 0$  implies  $0 \in \mathcal{A}_0$  and essinf  $\{X \in L_0^{\infty} \mid X \in \mathcal{A}_0\} = 0$ . Conversely, any set  $\mathcal{A}_0$  with these properties defines via (7) a conditional convex risk measure  $\rho_0$ .

Under an additional continuity condition, the conditional convex risk measure  $\rho_0$  admits the following dual representation in terms of suitably penalized probability measures  $Q \in \mathcal{M}_1(P)$ ; this is also called the *robust representation* of  $\rho_0$ .

For any  $Q \in \mathcal{M}_1(P)$  we define

$$\alpha_0(Q) := \operatorname{ess\,sup}_{X \in \mathcal{A}_0} E_Q\left[-X|\mathcal{F}_0\right].$$
(8)

Q-almost surely, taking the essential supremum under Q. Clearly,  $\alpha_0(Q)$  is well defined P-almost surely if Q is equivalent to P on  $\mathcal{F}_0$ , and in that case (8) can be read as well as an essential supremum under P.

- **Remark 2.** 1. Since  $0 \in A_0$ , we have  $\alpha_0(Q) \ge 0$  Q-a.s., and hence P-a.s. if  $Q \approx P$  on  $\mathcal{F}_0$ .
  - 2. For any  $X \in L^{\infty}$  we have  $X + \rho_0(X) \in \mathcal{A}_0$ , and so (8) implies

$$\rho_0(X) \ge E_Q \left[-X|\mathcal{F}_0\right] - \alpha_0(Q) \quad Q\text{-}a.s. \tag{9}$$

for any  $Q \in \mathcal{M}_1(P)$ .

With this definition of the penalty function  $\alpha_0$  the following equivalence holds; see [13, 4, 17, 6, 8, 1], and [19].

**Theorem 1.** For a conditional convex risk measure  $\rho_0$  with respect to  $\mathcal{F}_0$ , the following are equivalent:

1.  $\rho_0$  has the robust representation

$$\rho_0(X) = \underset{\substack{Q \in \mathcal{M}_1(P)\\Q \approx P \text{ on } \mathcal{F}_0}}{\operatorname{ess \, sup}} (E_Q [-X|\mathcal{F}_0] - \alpha_0(Q)), \quad X \in L^{\infty}, \quad (10)$$

where the essential supremum is taken under P.

2.  $\rho_0$  is continuous from above, i.e.,

$$X_k \searrow X$$
  $P\text{-}a.s \implies \rho_0(X_k) \nearrow \rho_0(X)$   $P\text{-}a.s$ 

for  $X \in L^{\infty}$  and any sequence  $(X_k) \subseteq L^{\infty}$ .

**Remark 3.** The penalty function  $\alpha_0$  is minimal in the following sense: If the representation (10) holds with some function  $\tilde{\alpha}_0$ , then

$$\tilde{\alpha}_0(Q) \ge \alpha_0(Q) \quad P\text{-}a.s. \tag{11}$$

for any  $Q \in \mathcal{M}_1(P)$  such that  $Q \approx P$  on  $\mathcal{F}_0$ . Indeed, (10) implies

$$\tilde{\alpha}_0(Q) \ge E_Q \left[-X|\mathcal{F}_0\right] - \rho_0(X) = E_Q \left[-(X+\rho_0(X))|\mathcal{F}_0\right] \quad P\text{-a.s.},$$

and hence (11) in view of (8), since  $X + \rho_0(X) \in \mathcal{A}_0$ .

**Remark 4.** Continuity from above is equivalent to the following condition, also called the Fatou property:

$$\rho_0(X) \le \liminf_{k \to \infty} \rho_0(X_k)$$

for any uniformly bounded sequence  $(X_k) \subset L^{\infty}$  which converges *P*-a.s. to some  $X \in L^{\infty}$ . We say that  $\rho_0$  has the Lebesgue property, if the inequality in the preceding condition can be replaced by the equality

$$\rho_0(X) = \lim_{k \to \infty} \rho_0(X_k).$$

The Lebesgue property holds if and only if  $\rho_0$  is not only continuous from above but also continuous from below, that is,

$$X_k \nearrow X \quad P\text{-}a.s \implies \rho_0(X_k) \searrow \rho_0(X) \quad P\text{-}a.s..$$

Moreover, it can be shown that the Lebesgue property is equivalent to the condition that the essential supremum in (10) is actually attained by some measure Q depending on X; for a proof in the unconditional case  $\mathcal{F}_0 = \{\Omega, \emptyset\}$  see [12, Theorem 2].

The proof of Theorem 1 shows that the robust representation in (10) can actually be refined in the sense that we can use a smaller set of probability measures; see, e.g., [17] or [19, Chapter 11].

**Corollary 1.** If  $\rho_0$  is continuous from above then we have

$$\rho_0(X) = \underset{Q \in \mathcal{Q}_0}{\operatorname{ess\,sup}} (E_Q \left[ -X | \mathcal{F}_0 \right] - \alpha_0(Q)), \quad X \in L^{\infty},$$

where

$$\mathcal{Q}_0 := \left\{ Q \in \mathcal{M}_1(P) \mid Q = P \text{ on } \mathcal{F}_0, \ E_Q[\alpha_0(Q)] < \infty \right\}.$$

**Remark 5.** In the special case  $\mathcal{F}_0 = \{\Omega, \emptyset\}$ , the preceding discussion reduces to standard definitions and basic facts for (unconditional) convex risk measures

$$\rho: L^{\infty} \to \mathbb{R}$$

on the Banach space  $L^{\infty}$ ; see, [2, 3, 10, 18, 20], and also [19, Chapter 4].

#### 2.2 Consistency

Let us now fix two sub- $\sigma$ -fields  $\mathcal{F}_0 \subseteq \mathcal{F}_1$  of  $\mathcal{F}$ . For i = 0, 1, we write  $L_i^{\infty} := L^{\infty}(\Omega, \mathcal{F}_i, P)$ , and we consider a conditional convex risk measure  $\rho_i : L^{\infty} \to L_i^{\infty}$  with respect to  $\mathcal{F}_i$ .

**Definition 2.** We say that the conditional risk measures  $\rho_0$  and  $\rho_1$  are consistent if

$$\rho_0 = \rho_0(-\rho_1),$$

that is, if  $\rho_0(-\rho_1(X)) = \rho_0(X)$  for all  $X \in L^{\infty}$ .

From now on we assume that both  $\rho_0$  and  $\rho_1$  are continuous from above. Let  $\mathcal{A}_i$  and  $\alpha_i$  denote the acceptance set and the minimal penalty function corresponding to  $\rho_i$ . Consistency of  $\rho_0$  and  $\rho_1$  can then be characterized in terms of the acceptance sets, in terms of the minimal penalty functions, and in terms of the joint behavior of  $(\rho_i)$  and  $(\alpha_i)$ . To this end, consider the restriction of  $\rho_0$  to the subspace  $L_1^{\infty}$  of  $L^{\infty}$  and denote by

$$\mathcal{A}_{0,1} := \left\{ X \in L_1^{\infty} \mid \rho_0(X) \le 0 \ P\text{-a.s.} \right\}$$

the acceptance set and by

$$\alpha_{0,1}(Q) := \operatorname{ess\,sup}_{X \in \mathcal{A}_{0,1}} E_Q\left[-X|\mathcal{F}_0\right], \qquad Q \in \mathcal{M}_1(P)$$

the minimal penalty function associated to this restriction in analogy to (8). As shown in [11, 17, 5, 8, 9, 6, 1], consistency can now be characterized as follows

Theorem 2. The following conditions are equivalent:

- 1.  $\rho_0$  and  $\rho_1$  are consistent.
- 2.  $A_0 = A_{0,1} + A_1$ .
- 3. For any  $Q \in \mathcal{M}_1(P)$ ,

$$\alpha_0(Q) = \alpha_{0,1}(Q) + E_Q[\alpha_1(Q) | \mathcal{F}_0] \quad Q\text{-}a.s.$$

4. For  $X \in L^{\infty}$  and any  $Q \in \mathcal{M}_1(P)$ ,

$$E_Q[\rho_1(X) + \alpha_1(Q) | \mathcal{F}_0] \le \rho_0(X) + \alpha_0(Q) \quad Q\text{-}a.s..$$

**Remark 6.** All our penalty functions are non-negative, since we have assumed that all our risk measures are normalized. Thus property (3) of Theorem 2 implies that

$$\alpha_0(Q) \ge E_Q[\alpha_1(Q)|\mathcal{F}_0] \quad Q\text{-a.s. for all } Q \in \mathcal{M}_1(P).$$
(12)

In particular,  $(\alpha_i(Q))_{i=0,1}$  is a non-negative supermartingale with respect to Q for all  $Q \in \mathcal{M}_1(P)$  such that  $E_Q[\alpha_0(Q)] < \infty$ . Note that the consistency criterion (3) of Theorem 2 provides, in addition to the supermartingale inequality (12), a special form of the predictable increasing process in the Doob decomposition of  $(\alpha_i)_{i=0,1}$ .

Condition (12) is equivalent to weak consistency of  $(\rho_i)_{i=0,1}$ , that is, to the condition that

$$\rho_1(X) \le 0 \implies \rho_0(X) \le 0$$

for any  $X \in L^{\infty}$ ; cf. [1, Proposition 8]. Note that weak consistency amounts to the relaxation  $\mathcal{A}_1 \subseteq \mathcal{A}_0$  of the consistency criterion (2) in Theorem 2. For other relaxations of the strong notion of consistency in Definition 2 see, for example, [26, 27, 25, 24, 1, 14], and in the law-invariant case [28].

In Section 4 we are going to use the Lebesgue property of conditional risk measures that was introduced in Remark 4, and we will apply the criterion of Proposition 1. This involves the following notion of strong sensitivity; see also [24].

**Definition 3.** We call a conditional convex risk measure  $\rho_0$  strongly sensitive with respect to P if

$$P[\rho_0(X) < \rho_0(Y)] > 0$$

whenever  $X, Y \in L^{\infty}$  satisfy  $X \ge Y$  *P*-a.s. and P[X > Y] > 0.

**Proposition 1.** Let  $\rho_0$  and  $\rho_1$  be consistent, and assume that  $\rho_0$  has the Lebesgue property and is strongly sensitive. Then  $\rho_1$  inherits the Lebesgue property and is strongly sensitive.

*Proof.* For  $X \in L^{\infty}$  and a uniformly bounded sequence  $(X_k)$  in  $L^{\infty}$  such that  $X_k \to X$  *P*-a.s., the Fatou property of  $\rho_1$  yields

$$\rho_1(X) \le \liminf_k \rho_1(X_k) \le \limsup_k \rho_1(X_k) \quad P\text{-a.s..}$$
(13)

To prove the Lebesgue property of  $\rho_1$ , we have to show that

$$\rho_1(X) = \limsup_k \rho_1(X_k) \quad P\text{-a.s.}.$$

In view of (13), this will follow from

$$\rho_0(-\rho_1(X)) = \rho_0(-\limsup_k \rho_1(X_k)),$$

due to the strong sensitivity of  $\rho_0$ . Indeed, using consistency, monotonicity of  $\rho_0$  applied to (13), and first the Fatou property and then the Lebesgue property of  $\rho_0$ , we obtain

$$\rho_0(X) = \rho_0(-\rho_1(X)) \le \rho_0(-\limsup_k \rho_1(X_k)) = \rho_0(\liminf_k \rho_1(-X_k)) \le \liminf_k \rho_0(-\rho_1(X_k)) = \liminf_k \rho_0(X_k) = \rho_0(X).$$

To see that  $\rho_1$  is strongly sensitive, take  $X, Y \in L^{\infty}$  such that  $X \ge Y$  and P[X > Y] > 0. Then we have  $P[\rho_1(X) < \rho_1(Y)] > 0$ , since  $\rho_1(X) = \rho_1(Y)$  *P*-a.s. would imply

$$\rho_0(X) = \rho_0(-\rho_1(X)) = \rho_0(-\rho_1(Y)) = \rho_0(Y)$$

in contradiction to the strong sensitivity of  $\rho_0$ .

### **3** Backwards Convergence

From now on we fix a filtration  $(\mathcal{F}_n)_{n\leq 0}$  on our probability space  $(\Omega, \mathcal{F}, P)$ . Thus, the  $\sigma$ -fields  $\mathcal{F}_n \subseteq \mathcal{F}$  are decreasing as n decreases to  $-\infty$ . We denote by

$$\mathcal{F}_{-\infty} := \bigcap_{n \le 0} \mathcal{F}_n$$

the corresponding *tail field* and write  $L_n^{\infty} = L^{\infty}(\Omega, \mathcal{F}_n, P)$ .

Let  $(\rho_n)_{n\leq 0}$  be a sequence of conditional convex risk measures

$$\rho_n: L^\infty \to L^\infty_n.$$

We denote by  $\mathcal{A}_n$  the acceptance set of  $\rho_n$ , and we assume that each  $\rho_n$  is continuous from above. Thus  $\rho_n$  admits a dual representation (10) in terms of its minimal penalty function  $\alpha_n$ . We also assume that the sequence is *consistent* in the sense that

$$\rho_n(-\rho_{n+1}) = \rho_n \tag{14}$$

for all n < 0.

**Example 1.** For  $\beta \ge 0$  consider the conditional entropic risk measures  $(\rho_n)_{n \le 0}$  defined by

$$\rho_n(X) := \frac{1}{\beta} \log E_P \left[ e^{-\beta X} | \mathcal{F}_n \right]; \tag{15}$$

for  $\beta = 0$  this is interpreted as the linear case (1), that is, as the limiting case of (15) as  $\beta$  decreases to 0. For  $\beta > 0$ , the corresponding penalty functions are given by

$$\alpha_n(Q) = \frac{1}{\beta} H_n(Q|P),$$

where  $H_n(Q|P)$  denotes the conditional relative entropy with respect to  $\mathcal{F}_n$ ; see [17] or [19, Chapter 11]. It is easy to check that the sequence  $(\rho_n)_{n\leq 0}$  is consistent. Note that  $\rho_n$  is law-invariant in the sense that  $\rho_n(X)$  only depends on the conditional distribution of X with respect to  $\mathcal{F}_n$  under P. Conversely, law-invariance together with consistency implies that the risk measures  $\rho_n$  are entropic, if the parameter  $\beta$  is allowed to be tail-measurable with values in  $[0, \infty)$ ; see [15] and also [22]. In this special entropic case, the sequence  $(\rho_n)_{n\leq 0}$  admits an immediate extension

$$\rho_{-\infty}(X) = \frac{1}{\beta} \log E_P \left[ e^{-\beta X} | \mathcal{F}_{-\infty} \right]$$

to the tail field, and the properties of asymptotic precision are clearly satisfied.

**Remark 7.** In general, let  $(\tilde{\rho_n})_{n \leq 0}$  be any sequence of conditional convex risk measures, not necessarily consistent. Defining recursively

$$\rho_0 := \tilde{\rho}_0 \quad and \quad \rho_n := \tilde{\rho_n}(-\rho_{n+1}) \quad for \ n < 0,$$

we obtain a sequence  $(\rho_n)_{n\leq 0}$  which is indeed consistent.

Our goal in this section is to extend the sequence  $(\rho_n)_{n\leq 0}$  to a conditional convex risk measure  $\rho_{-\infty}$  with respect to the tail field, to show that this risk measure is continuous from above, and to identify its dual representation. To this end we will make use of the supermartingale properties implied by the consistency condition (14), as they are stated in Theorem 2 and Remark 6.

Theorem 3. Let us assume

$$\sup_{n \le 0} E_P[\alpha_n(P)] < \infty.$$
(16)

Then the limit

$$\rho_{-\infty}(X) := \lim_{n \downarrow -\infty} \rho_n(X)$$

exists P-a.s. and in  $L^1(P)$  for all  $X \in L^{\infty}$ . Moreover, the resulting map

$$\rho_{-\infty}: L^{\infty} \to L^{\infty}(\Omega, \mathcal{F}_{-\infty}, P)$$

defines a conditional convex risk measure with respect to the tail-field  $\mathcal{F}_{-\infty}$ , and it satisfies the consistency condition

$$\rho_{-\infty} = \rho_{-\infty}(-\rho_n) \tag{17}$$

for all  $n \leq 0$ .

*Proof.* Fix  $X \in L^{\infty}$ . Due to our assumption (16), Theorem 2 together with Remark 6 shows that  $(\alpha_n(P))_{n\leq 0}$  is a backwards supermartingale under P which is bounded in  $L^1(\Omega, \mathcal{F}, P)$ . In view of part (4) of Theorem 2, the same is true for the process

$$V_n(P,X) := \rho_n(X) + \alpha_n(P), \qquad n \le 0,$$

since it is bounded from below by  $-\|X\|_{\infty}$  and satisfies

$$\sup_{n \le 0} E_P\left[V_n(P, X)\right] \le \|X\|_{\infty} + \sup_{n \le 0} E_P\left[\alpha_n(P)\right] < \infty.$$

Applying supermartingale convergence backwards under P, we obtain the existence of finite limits

$$V_{-\infty}(P,X) := \lim_{n \downarrow -\infty} V_n(P,X)$$
(18)

and

$$\alpha_{-\infty}(P) := \lim_{n \downarrow -\infty} \alpha_n(P)$$

both P-a.s. and in  $L^1(P)$ ; cf. [23]. This yields the existence of the limit

$$\rho_{-\infty}(X) := \lim_{n \downarrow -\infty} \rho_n(X) = V_{-\infty}(P, X) - \alpha_{-\infty}(P)$$
(19)

both P-a.s. and in  $L^1(P)$ . Moreover, we have  $|\rho_{-\infty}(X)| \leq ||X||_{\infty}$ , and it is easy to check that the resulting map

$$\rho_{-\infty}: L^{\infty} \to L^{\infty}(\Omega, \mathcal{F}_{-\infty}, P)$$

has the properties of a conditional convex risk measure with respect to the tail field  $\mathcal{F}_{-\infty}$ , as stated in Definition 1. To prove the consistency property (17) of  $\rho_{-\infty}$ , note that property (14) of the sequence  $(\rho_n)$  implies

$$\rho_{-\infty}(-\rho_n(X)) = \lim_{m \downarrow -\infty} \rho_m(-\rho_n(X)) = \lim_{m \downarrow -\infty} \rho_m(X) = \rho_{-\infty}(X)$$
  
y  $X \in L^{\infty}$  and  $n < 0.$ 

for any  $X \in L^{\infty}$  and  $n \leq 0$ .

In the preceding proof, we can replace the reference measure P by any measure Q belonging to the set

$$\mathcal{Q}_P := \left\{ Q \in \mathcal{M}_1(P) \mid Q = P \text{ on } \mathcal{F}_{-\infty}, \sup_{n \le 0} E_Q[\alpha_n(Q)] < \infty \right\}.$$

This yields the following result.

**Corollary 2.** For any  $Q \in \mathcal{Q}_P$ , the limit

$$\alpha_{-\infty}(Q) := \lim_{n \downarrow -\infty} \alpha_n(Q) \tag{20}$$

exists Q-a.s and in  $L^1(Q)$ , and we have

$$E_Q[\alpha_{-\infty}(Q)] = \lim_{n \downarrow -\infty} E_Q[\alpha_n(Q)] < \infty.$$

Let us denote by

$$\alpha_{n,n+1}(Q) := \operatorname{ess\,sup}_{X \in \mathcal{A}_n \cap L_{n+1}^{\infty}} E_Q\left[-X|\mathcal{F}_n\right]$$

the one-step penalty function of  $Q \in \mathcal{M}_1(P)$  for  $n \leq 0$ ; we put  $L_1^{\infty} := L^{\infty}$  so that  $\alpha_{0,1}(Q) = \alpha_0(Q)$ .

**Lemma 1.** For any  $Q \in \mathcal{Q}_P$  the limit  $\alpha_{-\infty}(Q)$  in (20) is given by

$$\alpha_{-\infty}(Q) = E_Q \left[ \sum_{l=-\infty}^{0} \alpha_{l,l+1}(Q) | \mathcal{F}_{-\infty} \right],$$
(21)

and we have

$$\alpha_{-\infty}(Q) = \lim_{n \downarrow -\infty} E_Q \left[ \alpha_n(Q) | \mathcal{F}_\infty \right]$$
(22)

Q-a.s. and in  $L^1(Q)$ .

*Proof.* Iterating condition (3) of Theorem 2 for l = n, ..., -1, we obtain

$$\alpha_n(Q) = \alpha_{n,n+1}(Q) + E_Q\left[\alpha_{n+1}(Q)|\mathcal{F}_n\right] = E_Q\left[\sum_{l=n}^0 \alpha_{l,l+1}(Q)|\mathcal{F}_n\right]$$
(23)

for any  $n \leq 0$ . Combining monotone convergence with martingale convergence ("Hunt's lemma"), we obtain

$$\alpha_{-\infty}(Q) = \lim_{n \downarrow -\infty} \alpha_n(Q) = E_Q \left[ \sum_{l=-\infty}^0 \alpha_{l,l+1}(Q) | \mathcal{F}_{-\infty} \right]$$

Q-a.s. and in  $L^1(Q)$ . Moreover, (23) implies

$$E_Q\left[\alpha_n(Q)|\mathcal{F}_{-\infty}\right] = E_Q\left[\sum_{l=n}^0 \alpha_{l,l+1}(Q)\Big|\mathcal{F}_{-\infty}\right],$$

and so equation (22) follows by monotone convergence.

Out next goal is to show that the conditional risk measure  $\rho_{-\infty}$  has the Fatou property, and that the minimal penalty function in its dual representation is given by the limits  $\alpha_{-\infty}(Q)$  for  $Q \in \mathcal{Q}_P$ . To this end we consider the functional  $\rho_P : L^{\infty} \to \mathbb{R}$  defined by

$$\rho_P(X) := E_P\left[\rho_{-\infty}(X)\right]. \tag{24}$$

**Lemma 2.**  $\rho_P$  is a convex risk measure, and it has the Fatou property.

*Proof.* It is easy to see that  $\rho_P$  has the properties of a convex risk measure. It remains to prove the Fatou property. Take  $X \in L^{\infty}$  and a uniformly bounded sequence  $(X_k)_{k \in \mathbb{N}}$  such that  $X_k \to X$  *P*-a.s.. For any  $n \leq 0$ , the Fatou property of  $\rho_n$  implies that the functional  $V_n(P, \cdot)$  has the Fatou property as well. Thus we obtain

$$E_P[V_n(P,X)] \le \liminf_k E_P[V_n(P,X_k)] \le \liminf_k E_P[V_{-\infty}(P,X_k)]$$

for all n, where the last inequality follows from the supermartingale property of  $(V_n(P, X_k))_{n \leq 0}$ . Using the supermartingale convergence in (18), this implies

$$E_P\left[V_{-\infty}(P,X)\right] \le \liminf_{n \to \infty} E_P\left[V_{-\infty}(P,X_k)\right].$$

Subtracting  $E_P[\alpha_{-\infty}(P)]$  from both sides and recalling (19), we obtain the Fatou property for  $\rho_P$ :

$$E_P\left[\rho_{-\infty}(X)\right] \le \liminf_k E_P\left[\rho_{-\infty}(X_k)\right].$$

**Theorem 4.** Under our assumptions (16) and (14), the conditional convex risk measure  $\rho_{-\infty}$  has the Fatou property, and it admits the representation

$$\rho_{-\infty}(X) = \operatorname{ess\,sup}_{Q \in \mathcal{Q}_P} \left( E_Q \left[ -X | \mathcal{F}_{-\infty} \right] - \alpha_{-\infty}(Q) \right), \quad X \in L^{\infty}$$

in terms of the limiting penalty function  $\alpha_{-\infty}$  in (20) and (21). Moreover,  $\alpha_{-\infty}$  coincides with the minimal penalty function of  $\rho_{-\infty}$ , i.e.,

$$\alpha_{-\infty}(Q) = \operatorname{ess\,sup}_{X \in \mathcal{A}_{-\infty}} E_Q \left[ -X | \mathcal{F}_{-\infty} \right] \quad P\text{-a.s.}$$
<sup>(25)</sup>

for any  $Q \in \mathcal{Q}_P$ , where

$$\mathcal{A}_{-\infty} := \left\{ X \in L^{\infty} \mid \rho_{-\infty}(X) \le 0 \right\}$$

denotes the acceptance set of  $\rho_{-\infty}$ .

*Proof.* To prove the Fatou property, we show that  $\rho_{-\infty}$  is continuous from above. Take  $X \in L^{\infty}$  and a decreasing sequence  $(X_k)$  in  $L^{\infty}$  with  $X_k \searrow X$  *P*-a.s.. Monotonicity of  $\rho_{-\infty}$  yields

$$\rho_{-\infty}(X) \ge \lim_{k} \rho_{-\infty}(X_k) \quad P\text{-a.s..}$$
(26)

On the other hand, the unconditional convex risk measure  $\rho_P$  in (24) is continuous from above by Lemma 2 and Remark 4. This implies

$$E_P \left[ \rho_{-\infty}(X) \right] = \rho_P(X) = \lim_k \rho_P(X_k) = \lim_k E_P \left[ \rho_{-\infty}(X_k) \right]$$
$$= E_P \left[ \lim_k \rho_{-\infty}(X_k) \right],$$

using monotone convergence in the last step. Combined with (26) this yields

$$\rho_{-\infty}(X) = \lim_{k} \rho_{-\infty}(X_k) \quad P\text{-a.s.},$$

and hence the Fatou property of  $\rho_{-\infty}$ . By Corollary 1 it follows that  $\rho_{-\infty}$  has the robust representation

$$\rho_{-\infty}(X) = \operatorname{ess\,sup}_{Q \in \mathcal{Q}_P} \left( E_Q \left[ -X | \mathcal{F}_{-\infty} \right] - \tilde{\alpha}_{-\infty}(Q) \right), \quad X \in L^{\infty},$$

where we denote by  $\tilde{\alpha}_{-\infty}(Q)$  the right-hand side of (25).

For a given  $Q \in \mathcal{Q}_P$ , we now show that  $\tilde{\alpha}_{-\infty}(Q) = \alpha_{-\infty}(Q)$  *P*-a.s.. Note first that, due to (17) and Theorem 2,

$$\tilde{\alpha}_{-\infty}(Q) = \tilde{\alpha}_{-\infty,n}(Q) + E_Q \left[ \alpha_n(Q) | \mathcal{F}_{-\infty} \right]$$
  

$$\geq E_Q \left[ \alpha_n(Q) | \mathcal{F}_{-\infty} \right] \qquad Q\text{-a.s.}$$

for any  $n \leq 0$ , where  $\tilde{\alpha}_{-\infty,n}(Q) \geq 0$  denotes the minimal penalty function of  $\rho_{-\infty}$  restricted to  $\mathcal{F}_n$ . This implies

$$\tilde{\alpha}_{-\infty}(Q) \ge \alpha_{-\infty}(Q)$$
 *P*-a.s., (27)

using equation (22) in Lemma 1 and the equality of Q and P on  $\mathcal{F}_{-\infty}$ . To obtain the converse inequality, take any  $X \in L^{\infty}$ . We have

$$\rho_n(X) \ge E_Q \left[-X|\mathcal{F}_n\right] - \alpha_n(Q) \quad Q\text{-a.s}$$

for any  $n \leq 0$  by Remark 2, and hence

$$\rho_{-\infty}(X) = \lim_{n} \rho_n(X) \ge \lim_{n} \left( E_Q \left[ -X | \mathcal{F}_n \right] - \alpha_n(Q) \right)$$
$$= E_Q \left[ -X | \mathcal{F}_{-\infty} \right] - \alpha_{-\infty}(Q) \quad Q\text{-a.s.}.$$

Since Q = P on  $\mathcal{F}_{-\infty}$ , we obtain

$$\rho_{-\infty}(X) \ge E_Q \left[-X|\mathcal{F}_{-\infty}\right] - \alpha_{-\infty}(Q) \quad P\text{-a.s..}$$
(28)

This holds for all  $Q \in \mathcal{Q}_P$ , and so we get

$$\rho_{-\infty}(X) = \operatorname{ess\,sup}_{Q \in \mathcal{Q}_P} \left( E_Q \left[ -X | \mathcal{F}_{-\infty} \right] - \tilde{\alpha}(Q)_{-\infty} \right) \\ \leq \operatorname{ess\,sup}_{Q \in \mathcal{Q}_P} \left( E_Q \left[ -X | \mathcal{F}_{-\infty} \right] - \alpha_{-\infty}(Q) \right) \\ \leq \rho_{-\infty}(X),$$

where we have used (27) for the first and (28) for the second inequality. The resulting equality shows that  $\rho_{-\infty}$  has a robust representation with penalty function  $\alpha_{-\infty}$ . Since  $\tilde{\alpha}_{-\infty}$  is the minimal penalty function, we obtain  $\tilde{\alpha}_{-\infty}(Q) \leq \alpha_{-\infty}(Q)$  *P*-a.s. for any  $Q \in Q_P$ . Combined with (27), this yields equality (25).

Thus we have shown backwards convergence of  $\rho_n$  as  $n \to -\infty$  to a nice conditional risk measure  $\rho_{-\infty}$  with respect to the tail field. This can be seen as a backward analogue to the properties of asymptotic safety and asymptotic precision in the forward direction for a consistent sequence  $(\rho_n)_{n\geq 0}$  along a filtration  $(\mathcal{F}_n)_{n\geq 0}$ ; see the discussion in [17, Section 5]. As shown in [17, Theorem 5.4], asymptotic safety can be characterized in terms of various asymptotic properties of acceptance sets and of penalty functions as n tends to  $\infty$ . The following corollary states backward analogues to those properties as n tends to  $-\infty$ .

For  $n \leq 0$ , we denote by

 $\mathcal{A}_{-\infty,n} := \mathcal{A}_{-\infty} \cap L_n^{\infty}$ 

the acceptance set of  $\rho_{-\infty}$  restricted to  $\mathcal{F}_n$ , and by

$$\alpha_{-\infty,n}(Q) := \operatorname{ess\,sup}_{X \in \mathcal{A}_{-\infty,n}} E_Q\left[-X|\mathcal{F}_n\right],$$

the corresponding minimal penalty function for  $Q \in Q_P$ .

Corollary 3. 1.  $\bigcap_n \mathcal{A}_{-\infty,n} = L^{\infty}_+(\mathcal{F}_{-\infty}).$ 

- 2.  $\lim_{n\downarrow-\infty} \alpha_{-\infty,n}(Q) = 0$  Q-a.s. for all  $Q \in \mathcal{Q}_P$ .
- (3) A position  $X \in L^{\infty}$  belongs to  $\mathcal{A}_{-\infty}$  if and only if there exists a uniformly bounded sequence  $X_n \in \mathcal{A}_n$ ,  $n \leq 0$ , such that  $\exists \lim_{n \downarrow -\infty} X_n \leq X$ .

*Proof.* Since each  $\rho_n$  and hence  $\rho_{-\infty}$  is normalized, we obtain  $L^{\infty}_+(\mathcal{F}_{-\infty}) \subseteq \mathcal{A}_{-\infty,n}$  for any  $n \leq 0$ , and this shows the inclusion " $\supseteq$ " in (1). Conversely, if  $X \in \mathcal{A}_{-\infty}$  is  $\mathcal{F}_n$ -measurable for all  $n \leq 0$ , then X is  $\mathcal{F}_{-\infty}$ -measurable, and conditional cash invariance of  $\rho_{-\infty}$  yields  $-X = \rho_{-\infty}(X) \leq 0$  *P*-a.s.. This proves (1).

By Theorem 2 combined with (25), the consistency relation

$$\rho_{-\infty}(-\rho_n) = \rho_{-\infty}$$

in Theorem 3 implies

$$\alpha_{-\infty}(Q) = \alpha_{-\infty,n}(Q) + E_Q \left[ \alpha_n(Q) | \mathcal{F}_{-\infty} \right]$$

for any  $Q \in \mathcal{Q}_P$ , and so the convergence in (2) follows from the second equality in Lemma 1.

In order to prove (3), take  $X \in \mathcal{A}_{-\infty}$ . Note that  $X_n := X + \rho_n(X) \in \mathcal{A}_n$  for all  $n \leq 0$ , that the sequence  $(X_n)$  is uniformly bounded by  $2||X||_{\infty}$ , and that  $\lim_n X_n = X + \rho_{-\infty}(X) \leq X$ . Conversely, let  $X \in L^{\infty}$  satisfy the condition  $X \geq \lim_n X_n$  for some uniformly bounded sequence  $(X_n)$  such that  $X_n \in \mathcal{A}_n$ for each  $n \leq 0$ . Since  $\rho_n(X_n) \leq 0$ , monotonicity and the Fatou property of  $\rho_{-\infty}$ together with the consistency condition  $\rho_{-\infty}(-\rho_n) = \rho_{-\infty}$  yield

$$\rho_{-\infty}(X) \le \rho_{-\infty}(\lim_{n} X_n) \le \liminf_{n} \rho_{-\infty}(X_n)$$
$$= \liminf_{n} \rho_{-\infty}(-\rho_n(X_n)) \le 0,$$

and so X belongs to  $\mathcal{A}_{-\infty}$ .

For the rest of this section we focus on the special case where each  $\rho_n$  is *coherent*. Let us denote by  $\mathcal{M}_1^e(P)$  the class of all probability measures Q on  $(\Omega, \mathcal{F})$  which are equivalent to P.

**Definition 4.** A class  $Q \subseteq \mathcal{M}_1^e(P)$  of probability measures on  $(\Omega, \mathcal{F})$  is called stable with respect to the filtration  $(\mathcal{F}_n)_{n\leq 0}$  if, for any  $Q_1, Q_2 \in \mathcal{Q}$  and any  $n \leq 0$ , the probability measure Q defined by

$$E_Q[X] := E_{Q_1}[E_{Q_2}[X|\mathcal{F}_n]]$$

belongs again to Q.

**Corollary 4.** Under assumptions (16) and (14), the conditional risk measure  $\rho_{-\infty}$  defined in Theorem 3 is coherent if and only if each  $\rho_n$ ,  $n \leq 0$ , is coherent. In this case both  $\rho_n$  and  $\rho_{-\infty}$  have robust representations in terms of the set

$$\mathcal{Q}_P^e := \mathcal{Q}_P \cap \mathcal{M}_1^e(P) = \left\{ Q \approx P \mid \alpha_{-\infty}(Q) = 0 \right\},$$

*i.e.*,

$$\rho_n(X) = \operatorname{ess\,sup}_{Q \in \mathcal{Q}_P^e} E_Q \left[ -X | \mathcal{F}_n \right], \qquad X \in L^{\infty}$$

for  $n \leq 0$  and  $n = -\infty$ . Moreover, the set  $\mathcal{Q}_P^e$  is stable, and the process  $(\rho_n(X))_{n\leq 0}$  is a backward Q-supermartingale for any  $Q \in \mathcal{Q}_P^e$ .

*Proof.* It is straightforward to see that the limiting risk measure  $\rho_{-\infty}$  is coherent, if each  $\rho_n$  is coherent. The converse as well as all other statements of the corollary follow from Theorem 4 and [17, Corollary 4.12], due to the consistency condition (17).

**Remark 8.** Under an additional condition of comonotonicity, a conditional version of the arguments in [19, Corollary 4.95] shows that the coherent risk measures  $\rho_n$  can be interpreted as conditional Choquet integrals

$$\rho_n(X) = \int (-X) dC_n,$$

that is, as Choquet integrals with respect to the conditional submodular capacity  $C_n$  defined by

$$C_n(A) := \rho_n(-I_A)$$

for any  $A \in \mathcal{F}$ ; see also [7] and the references therein. If these conditional Choquet integrals are consistent in the sense of (14), then Theorem 4 shows that they converge along our decreasing  $\sigma$ -fields, that is,

$$\lim_{n} \int X dC_n = \int X dC_{-\infty} \qquad P-a.s.$$

for any  $X \in L^{\infty}$ .

## 4 The structure of global risk measures consistent with $(\rho_n)_{n \leq 0}$

Let us denote by  $\mathcal{R}$  the class of all convex risk measures  $\rho$  on  $L^{\infty}$  which are consistent with the sequence  $(\rho_n)_{n\leq 0}$ , that is,  $\rho$  satisfies the condition

$$\rho = \rho(-\rho_n), \qquad n \le 0. \tag{29}$$

Note that  $\mathcal{R} \neq \emptyset$ , since the risk measure  $\rho_P$  defined in (24) belongs to  $\mathcal{R}$ .

From now on we focus on risk measures  $\rho \in \mathcal{R}$  which have the Lebesgue property. We denote be  $\mathcal{R}_L$  the class of all those risk measures, and by  $\mathcal{R}_{L,S}$  the subclass of all  $\rho \in \mathcal{R}_L$  which are strongly sensitive in the sense of Definition 3. In this section, our aim is to clarify the structure of the sets  $\mathcal{R}_L$  and  $\mathcal{R}_{L,S}$ .

**Lemma 3.** For any  $\rho \in \mathcal{R}_L$  the consistency condition (29) extends to the tail field  $\mathcal{F}_{-\infty}$ , that is,

$$\rho(-\rho_{-\infty}) = \rho.$$

*Proof.* For any  $X \in L^{\infty}$  the sequence  $(\rho_n(X))_{n\leq 0}$  is uniformly bounded by  $||X||_{\infty}$  and *P*-a.s. convergent to  $\rho_{-\infty}(X)$ . Combining (29) with the Lebesgue property of  $\rho$ , we obtain

$$\rho(-\rho_{-\infty}(X)) = \lim_{n} \rho(-\rho_n(X)) = \rho(X).$$
(30)

**Proposition 2.** We have  $\mathcal{R}_{L,S} \neq \emptyset$  if and only if the conditional risk measure  $\rho_{-\infty}$  has the Lebesgue property and is strongly sensitive in the sense of Definition 3.

*Proof.* Suppose that  $\mathcal{R}_{L,S} \neq \emptyset$ . For any  $\rho \in \mathcal{R}_{L,S}$ , we have  $\rho(-\rho_{-\infty}) = \rho$  due to Lemma 3. Applying Proposition 1 to  $\rho$  and to  $\mathcal{F}_1 := \mathcal{F}_{-\infty}$  we see that  $\rho_{-\infty}$  has the Lebesgue property and is strongly sensitive.

Conversely, the Lebesgue property of  $\rho_{-\infty}$  implies that the risk measure  $\rho_P$  defined in (24) belongs to  $\mathcal{R}_L$ . If, moreover,  $\rho_{-\infty}$  is strongly sensitive in the sense of Definition 3, then  $\rho := \rho_P$  is strongly sensitive as well.

Our description of the risk measures in  $\mathcal{R}_L$  and  $\mathcal{R}_{L,S}$  will involve the conditional risk measure  $\rho_{-\infty}$  with respect to the tail filed  $\mathcal{F}_{-\infty}$  and an unconditional risk measure on the tail field. More precisely, let us denote by  $\hat{\mathcal{R}}$  the class of convex risk measures  $\hat{\rho}$  on  $\hat{L}^{\infty} := L^{\infty}(\Omega, \mathcal{F}_{-\infty}, P)$  which have the Lebesgue property on  $\hat{L}^{\infty}$ , and by  $\hat{\mathcal{R}}_{L,S}$  the subclass of all  $\hat{\rho} \in \hat{\mathcal{R}}_L$  which are strongly sensitive on  $\hat{L}^{\infty}$ .

**Theorem 5.** Suppose that  $\rho_{-\infty}$  has the Lebesgue property. Then the class  $R_L$  has the following structure:

$$\mathcal{R}_L = \left\{ \hat{\rho}(-\rho_{-\infty}) \mid \hat{\rho} \in \hat{\mathcal{R}}_L \right\}.$$
(31)

*Proof.* Take any  $\rho \in \mathcal{R}_L$ , and denote by  $\hat{\rho}$  the restriction of  $\rho$  to  $\hat{L}^{\infty}$ . Clearly,  $\hat{\rho}$  belongs to  $\hat{\mathcal{R}}_L$ , and Lemma 3 implies

$$\rho = \rho(-\rho_{-\infty}) = \hat{\rho}(-\rho_{-\infty}).$$

This shows the inclusion " $\subseteq$ " in (31).

Conversely, take any  $\hat{\rho} \in \hat{\mathcal{R}}_L$ . Then  $\rho := \hat{\rho}(-\rho_{-\infty})$  defines a convex risk measure on  $L^{\infty}$ . We have  $\rho \in \mathcal{R}$  since

$$\rho(-\rho_n(X)) = \hat{\rho}(-\rho_{-\infty}(-\rho_n(X))) = \hat{\rho}(-\rho_{-\infty}(X)) = \rho(X)$$

due to (17). Moreover,  $\rho$  has the Lebesgue property on  $L^{\infty}$ . Indeed, for any uniformly bounded sequence  $(X_k)$  in  $L^{\infty}$  such that  $X_k \to X$  P-a.s., the Lebesgue property of  $\rho_{-\infty}$  implies

$$\rho_{-\infty}(X) = \lim_{k} \rho_{-\infty}(X_k) \quad P\text{-a.s.},$$

hence

$$\rho(X) = \hat{\rho}(-\rho_{-\infty}(X)) = \lim_{k} \hat{\rho}(-\rho_{-\infty}(X_k)) = \lim_{k} \rho(X_k)$$

due to the Lebesgue property of  $\hat{\rho}.$ 

Corollary 5. If  $\mathcal{R}_{L,S} \neq \emptyset$ , then

$$\mathcal{R}_{L,S} = \left\{ \hat{\rho}(-\rho_{-\infty}) \mid \hat{\rho} \in \hat{\mathcal{R}}_{L,S} \right\}.$$

*Proof.* By Proposition 2, the existence of some  $\rho \in \mathcal{R}_{L,S}$  implies that  $\rho_{-\infty}$  has the Lebesgue property and is strongly sensitive. For any  $\rho \in \mathcal{R}_{L,S}$ , the restriction  $\hat{\rho}$  of  $\rho$  to  $\hat{L}^{\infty}$  clearly belongs to  $\hat{\mathcal{R}}_{L,S}$ , and we have  $\rho = \hat{\rho}(-\rho_{-\infty})$  due to (30).

Conversely, take  $\hat{\rho} \in \hat{\mathcal{R}}_{L,S}$ . Then Theorem 5 shows that  $\rho := \hat{\rho}(-\rho_{-\infty})$ belongs to  $\mathcal{R}_L$ . Moreover,  $\rho$  is strongly sensitive. Indeed, for X and Y in  $L^{\infty}$ with  $X \leq Y$  P-a.s.and P[X < Y] > 0, we obtain  $\rho_{-\infty}(X) \geq \rho_{-\infty}(Y)$  P-a.s.and  $P[\rho_{-\infty}(X) > \rho_{-\infty}(Y)] > 0$  due to Proposition 2, and so the strong sensitivity of  $\hat{\rho}$  implies

$$\rho(X) = \hat{\rho}(-\rho_{-\infty}(X)) > \hat{\rho}(-\rho_{-\infty}(Y)) = \rho(Y).$$

**Remark 9.** Throughout this paper, we have worked with a fixed probability measure P. In the spatial setting of [16], the single measure P will be replaced by a whole class P of probability measures, namely the class of Gibbs measures with given local conditional probabilities. This will require a refined analysis, where conditional risk measures with respect to P are replaced by risk kernels, and where the results of the present paper will be used as building blocks.

### Acknowledgments

Convex risk measures contain, as a special case, the class of Choquet integrals with respect to a submodular Choquet capacity. In this sense they are related to probabilistic potential theory; see, e.g., [21] for a discussion of Choquet capacities in this context. Masatoshi Fukushima and the first author were both working in this area at the time when they first met at the Sixth Berkeley Symposium on Mathematical Statistics and Probability in 1970. It is a great pleasure to dedicate this paper to Masatoshi Fukushima on the occasion of his 80th birthday.

### References

- Beatrice Acciaio and Irina Penner. Dynamic risk measures. In G. Di Nunno and B. Øksendal (Eds.) Advanced Mathematical Methods for Finance, pages 1–34. Springer, Berlin Heidelberg, 2011.
- [2] Philippe Artzner, Freddy Delbaen, Jean-Marc Eber, and David Heath. Thinking coherently. *RISK*, 10(November):68–71, 1997.
- [3] Philippe Artzner, Freddy Delbaen, Jean-Marc Eber, and David Heath. Coherent measures of risk. *Math. Finance*, 9(3):203–228, 1999.
- [4] Jocelyne Bion-Nadal. Conditional risk measure and robust representation conditional of convex risk measures. 2004. CMAP preprint Ecole Polytechnique Palaiseau. 557,http://www.cmap.polytechnique.fr/preprint/repository/557.pdf.
- [5] Jocelyne Bion-Nadal. Dynamic risk measures: time consistency and risk measures from BMO martingales. *Finance Stoch.*, 12(2):219–244, 2008.
- [6] Jocelyne Bion-Nadal. Time consistent dynamic risk processes. Stochastic Processes and their Applications, 119:633–654, 2008.
- [7] Alain Chateauneuf, Robert Kast, and André Lapied. Conditioning capacities and Choquet integrals: the role of comonotony. *Theory and Decision*, 51(2-4):367–386 (2002), 2001. FUR X Conference (Turin, 2001).
- [8] Patrick Cheridito, Freddy Delbaen, and Michael Kupper. Dynamic monetary risk measures for bounded discrete-time processes. *Electron. J. Probab.*, 11:no. 3, 57–106 (electronic), 2006.
- [9] Patrick Cheridito and Michael Kupper. Composition of time-consistent dynamic monetary risk measures in discrete time. *International Journal of Theoretical and Applied Finance*, 14(1):137–162, 2011.
- [10] Freddy Delbaen. Coherent risk measures on general probability spaces. In K. Sandmann and P.J. Schönbucher (Eds.) Advances in Finance and Stochastics, pages 1–37. Springer, Berlin, 2002.

- [11] Freddy Delbaen. The structure of m-stable sets and in particular of the set of risk neutral measures. In In memoriam Paul-André Meyer: Séminaire de Probabilités XXXIX, volume 1874 of Lecture Notes in Math., pages 215– 258. Springer, Berlin, 2006.
- [12] Freddy Delbaen. Differentiability properties of utility functions. In Optimality and risk—modern trends in mathematical finance, pages 39–48. Springer, Berlin, 2009.
- [13] Kai Detlefsen and Giacomo Scandolo. Conditional and dynamic convex risk measures. *Finance Stoch.*, 9(4):539–561, 2005.
- [14] Vicky Fasen and Adela Svejda. Time consistency of multi-period distortion measures. Stat. Risk Model., 29(2):133–153, 2012.
- [15] Hans Föllmer. Spatial risk measures and their local specification: the locally law-invariant case. Statistics & Risk Modeling, 31(1):79–103, 2014.
- [16] Hans Föllmer and Claudia Klüppelberg. Spatial risk measures: local specification and boundary risk. *forthcoming*.
- [17] Hans Föllmer and Irina Penner. Convex risk measures and the dynamics of their penalty functions. *Statist. Decisions*, 24(1):61–96, 2006.
- [18] Hans Föllmer and Alexander Schied. Convex measures of risk and trading constraints. *Finance Stoch.*, 6(4):429–447, 2002.
- [19] Hans Föllmer and Alexander Schied. Stochastic finance: An introduction in discrete time. Walter de Gruyter & Co., Berlin, 3rd revised and extended edition, 2011.
- [20] Marco Frittelli and Emanuela Rosazza Gianin. Putting order in risk measures. Journal of Banking Finance, 26(7):1473–1486, 2002.
- [21] Masatoshi Fukushima. Dirichlet forms and Markov processes, volume 23 of North-Holland Mathematical Library. North-Holland Publishing Co., Amsterdam, 1980.
- [22] Michael Kupper and Walter Schachermayer. Representation results for law invariant time consistent functions. *Mathematics and Financial Economics*, 2(3):189–210, 2009.
- [23] J. Neveu. Discrete-parameter martingales. Translated by T. P. Speed. North-Holland Mathematical Library. Vol. 10. Amsterdam - Oxford: North- Holland Publishing Company; New York: American Elsevier Publishing Company, Inc. VIII, 236 p. Dfl. 62.50 \$ 26.95., 1975.
- [24] Berend Roorda and J. M. Schumacher. Membership conditions for consistent families of monetary valuations. To appear in Statistics & Risk Modelling. 10.1524/strm.2013.1131.

- [25] Berend Roorda and J. M. Schumacher. Time consistency conditions for acceptability measures, with an application to Tail Value at Risk. *Insurance Math. Econom.*, 40(2):209–230, 2007.
- [26] Sina Tutsch. Konsistente und konsequente dynamische Risikomaße und das Problem der Aktualisierung. PhD thesis, Humboldt-Universität zu Berlin, 2006.
- [27] Sina Tutsch. Update rules for convex risk measures. *Quant. Finance*, 8(8):833–843, 2008.
- [28] Stefan Weber. Distribution-invariant risk measures, information, and dynamic consistency. *Mathematical Finance*, 16(2):419–441, 2006.